# Quantifying child mortality reductions related to measles vaccination 

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## Supporting Information

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## Supporting Information S1. Data description and coding

Data used in the analyses were derived from multiple sources (Table S1A). Measles deaths were identified from the WHO Mortality Database based on ICD codes shown in Table S1B. Measles vaccine coverage estimates for years prior to 1980 were derived from a variety of country-specific reports shown in Table S1C. The countries, numbers of observations, and observation years with data present for all variables are shown in Table S1D.

Table S1A. Data sources and years for analyses

| Variables | Years | Source/Description |
| :--- | :--- | :--- |
| Cause/age-specific mortality counts <br> Size of age-specific population at risk <br> Age-specific all other cause death rate | $1960-2005$ | WHO Mortality Database (For <br> measles cause-specific ICD <br> codes used see Table S1B) |
| Quality rating for each country's vital registration <br> data | $\mathrm{n} / \mathrm{a}$ | Mathers CD, Fat DM, Inoue M, <br> Rao C, Lopez AD (2005) <br> Counting the dead and what they <br> died from: an assessment of the <br> global status of cause of death <br> data. Bull World Health Organ <br> $83: 171-177$. |
| Vaccination coverage | $1980-2005$ | WHO/UNICEF |
| Vaccination coverage and/or start date of <br> vaccination | $1960-1980$ | Various Source on Vaccination <br> (Complete list of sources in |
| Tecond dose of MCV used in vaccination | $1960-2005$ | WHO S1C) |
| Crude birth rate <br> Percent urban <br> Population size and structure <br> Population density | $1950-2005$ | United Nations World <br> Population Prospects |
| Real per-capita GDP (Laspeyres, Constant I\$2000) | $1950-2004$ | Penn World Tables 6.2 |

Abbreviations: ICD, International Classification of Diseases; MCV, measles-containing vaccine; WHO, World Health Organization; GDP, gross domestic product

| Table S1B. International Classification of Disease (ICD) codes used to identify measles deaths* |  |  |
| :--- | :--- | :--- |
| ICD Edition | Detailed List Numbers | Description |
| ICD-7 | 085 | Measles |
| ICD-8 | 055 | Measles |
| ICD-9 | 055 | Measles |
| ICD-10 | B05 | Measles |
| ICD-10 | B050 | Measles complicated by encephalitis |
| ICD-10 | B051 | Measles complicated by meningitis |
| ICD-10 | B052 | Measles complicated by pneumonia |
| ICD-10 | B053 | Measles complicated by otitis media |
| ICD-10 | B054 | Measles with intestinal complications |
| ICD-10 | B058 | Measles with other complications |
| ICD-10 | B059 | Measles without complication |

* Summary codes corresponding to these detailed list numbers were used to extract country, year, and genderspecific mortality counts for measles-related deaths.

Table S1C. Sources for Measles Vaccine Coverage prior to 1980

## Sources

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November 13]; Available from: http://www.euvac.net/graphics/euvac/vaccination/belgium.html.

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- Measles elimination plan: Spain (Instituto de Salud Carlos III). Instituto de Salud Carlos III; [cited 2007 November 13]; Available from: http://www.euvac.net/graphics/euvac/pdf/plan_spain.pdf.
- The Swiss Childhood Vaccine Schedule (EUVAC NET). EUVAC NET; [cited 2007 November 13]; Available from: http://www.euvac.net/graphics/euvac/vaccination/switzerland.html.
- Expanded Program on Immunization in Brazil: vaccination coverage. EPI Newsletter. 1979 December;1(4).
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- Epidemic outbreak of measles in three central provinces of Chile. EPI Newsletter. 1981;3(2).
- Panama: summary of immunization data from 1979 Family Planning/Maternal Child Health Survey. EPI Newsletter. 1981;3(3).
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Table S1D. Observations and years by country

| Country | Number of observations | First year | Last year |
| :---: | :---: | :---: | :---: |
| Austria | 37 | 1960 | 2004 |
| Azerbaijan | 8 | 1995 | 2002 |
| Belarus | 8 | 1996 | 2003 |
| Belgium | 31 | 1960 | 1997 |
| Belize | 9 | 1982 | 1996 |
| Brazil | 21 | 1979 | 1999 |
| Bulgaria | 13 | 1992 | 2004 |
| Canada | 17 | 1960 | 2000 |
| Chile | 28 | 1960 | 1996 |
| Colombia | 20 | 1960 | 1998 |
| Costa Rica | 28 | 1960 | 1996 |
| Cuba | 22 | 1979 | 2000 |
| Denmark | 27 | 1960 | 1994 |
| El Salvador | 14 | 1960 | 1997 |
| Finland | 34 | 1960 | 1995 |
| France | 35 | 1960 | 2003 |
| Germany | 13 | 1992 | 2004 |
| Guatemala | 24 | 1960 | 2003 |
| Hungary | 26 | 1971 | 1996 |
| Ireland | 42 | 1960 | 2005 |
| Israel | 16 | 1982 | 1997 |
| Italy | 39 | 1960 | 2002 |
| Kazakhstan | 5 | 1994 | 2004 |
| Kuwait | 8 | 1981 | 1994 |
| Kyrgyzstan | 6 | 1994 | 1999 |
| Luxembourg | 14 | 1984 | 1997 |
| Mexico | 36 | 1960 | 1997 |
| Netherlands | 37 | 1960 | 1999 |
| Norway | 32 | 1960 | 1995 |
| Panama | 18 | 1960 | 1997 |
| Republic of Korea | 16 | 1987 | 2002 |
| Romania | 33 | 1961 | 1998 |
| Russian Federation | 11 | 1994 | 2004 |
| Spain | 41 | 1960 | 2004 |
| Sweden | 27 | 1960 | 1996 |
| Switzerland | 34 | 1960 | 2004 |
| TFYR Macedonia | 9 | 1995 | 2003 |
| Turkmenistan | 5 | 1994 | 1998 |
| Ukraine | 11 | 1994 | 2004 |
| United Kingdom | 43 | 1960 | 2004 |
| United States of America | 40 | 1960 | 2001 |
| Uruguay | 16 | 1960 | 1996 |
| Uzbekistan | 8 | 1994 | 2003 |
| Venezuela | 18 | 1960 | 1994 |

## Supporting Information S2. Variable construction

## Logarithmic transformation of measles death rates

For analyses of reductions in measles death rates, the rates were log transformed (natural log). For country-years with zero observed measles deaths, the log-transformed rate is undefined. To prevent these observations from being dropped from the analysis, we therefore replaced zero values with the minimum observed rate divided by 10 and then log-transformed all rates. The minimum non-zero observed rate in the dataset occurred in the United States in 1992 and was $5 \times 10^{-3}$ measles deaths per 100,000 children aged 0 through 5 . Sensitivity analyses relating to the treatment of zeros and the analysis of measles specific death rates are presented in Supporting Information S6.

## Categorizing measles vaccination coverage levels

Measles-containing vaccine (MCV) coverage is a continuous variable running from 0 to $100 \%$. Because the relationship between MCV coverage and measles death rates may be non-linear and because we did not wish to impose a functional form, we categorized MCV coverage into a number of discrete levels. We defined these divisions prospectively so that the number of observations in each level above $0 \% \mathrm{MCV}$ coverage was nearly equal and cutoffs were divisible by 5 . MCV coverage was categorized into the following levels: $0 \% ; 1-59 \% ; 60-79 \% ; 80-89 \% ; 90-94 \%$; and $>=95 \%$ coverage. We also constructed restricted cubic splines for MCV coverage with knots placed at the same cutoffs [See: Luke Keele. Semiparametric Regression for Social Sciences. Chichester, England: John Wiley and Sons, Ltd. 2008].

## Supporting Information S3. Sensitivity analysis: model fit and specifications

We compared alternative model specifications using both the Akaike Information Criterion (AIC) and the Bayesian Information Criterion (BIC). In general a reduction in the AIC or BIC from one model specification to another of approximately 2 or greater indicates a significant improvement in model fit even after penalization for specifications that include additional parameters.

We compared the following model specifications for MCV coverage: 1) the base case model with all covariates but omitting any MCV coverage term (i.e., no MCV coverage term); 2) the base case model with all covariates and MCV coverage a continuous variable ranging from $0-1$ with one regression coefficient estimated for this term (i.e., linear MCV coverage term); 3) the base case model with all covariates and MCV coverage categorized into separate indicators for each coverage range as described in Supporting Information S2 (i.e., indicators MCV coverage terms); 4) the base case model with all covariates and MCV coverage entered as a series of restricted cubic splines as described in Supporting Information S2 (i.e., restricted cubic splines for MCV coverage). The model specification where MCV coverage enters as restricted cubic splines is preferred as it minimizes both the Akaike Information Criterion (AIC) and the Bayesian Information Criterion (BIC) (Table S3A).

Table S3A. Specification of MCV coverage in the model

| Specification | AIC | BIC |
| :--- | :--- | :--- |
| Model (no MCV coverage term) | 4252.497 | 4306.261 |
| Model (linear MCV coverage term) | 4205.153 | 4263.804 |
| Model (indicator MCV coverage terms) | 4202.318 | 4280.519 |
| Model (restricted cubic splines for MCV coverage) | 4201.080 | 4279.281 |

In general, the categorization of MCV coverage that was prospectively defined and used in the base case ( $1-60 \%, 60-80 \%, 80-90 \%, 90-95 \%, 95-100 \%$ ) also performed quite well. However, for the particular data used, several changes to the exact division (shown in italics in Table S3B) performed better in terms of minimizing the AIC and BIC. For example, changing the $80 \%$ cutoff to $70 \%$ or $75 \%$ and/or changing the $90 \%$ cutoff to $85 \%$ all improved both the AIC and BIC.

Table S3B. Alternate categorization of MCV coverage

| Specification | AIC | BIC |
| :--- | :--- | :--- |
| Model (no MCV coverage term) | 4252.497 | 4306.261 |
|  |  |  |
| Model (indicator MCV coverage, 1-60, 60-80, 80-90, 90-95, 95-100) | 4202.318 | 4280.519 |
| Model (indicator MCV coverage, 1-20, 20-40, 40-60, 60-80, 80-100) | 4204.893 | 4283.094 |
|  |  |  |
| Model (indicator MCV coverage, 1-10, 10-80, 80-90, 90-95, 95-100) | 4225.608 | 4303.809 |
| Model (indicator MCV coverage, 1-20, 20-80, 80-90, 90-95, 95-100) | 4218.788 | 4296.988 |
| Model (indicator MCV coverage, 1-30, 30-80, 80-90, 90-95, 95-100) | 4217.276 | 4295.477 |
| Model (indicator MCV coverage, 1-40, 40-80, 80-90, 90-95, 95-100) | 4216.87 | 4295.07 |
| Model (indicator MCV coverage, 1-50, 50-80, 80-90, 90-95, 95-100) | 4212.321 | 4290.522 |
|  |  |  |
| Model (indicator MCV coverage, 1-60, 60-70, 70-90, 90-95, 95-100) | 4207.591 | 4285.792 |
| Model (indicator MCV coverage, 1-60, 60-70, 70-80, 80-90, 90-100) | 4203.739 | 4281.939 |
| Model (indicator MCV coverage, 1-60, 60-70, 70-85, 85-95, 95-100) | 4201.444 | 4279.645 |
| Model (indicator MCV coverage, 1-60, 60-80, 80-85, 85-95, 95-100) | 4200.135 | 4278.336 |
| Model (indicator MCV coverage, $1-60,60-75,75-85,85-95,95-100$ ) | 4197.825 | 4276.026 |

However, even with the categorization of MCV coverage that produced the biggest improvement in AIC/BIC compared to the pre-specified categorization, the direction and significance of all regression coefficients remain largely unchanged compared to the base case. Table S3C shows results of this comparison for the linear regression of logged under-5 measles mortality as a function of categorical coverage indicators and control variables.

Table S3C. Comparison of regression coefficients under alternate categorization of MCV coverage levels*

| Best Specification (AIC/BIC) |  | Base Case Specification |  |  |  |
| :--- | :---: | :---: | :--- | :---: | :---: |
| Independent variables | Coefficient | P-value | Independent variables | Coefficient | P-value |
| MCV coverage of $1-59 \%$ | -0.240 | 0.514 | MCV coverage of $1-59 \%$ | -0.236 | 0.523 |
| MCV coverage of $60-74 \%$ | $\mathbf{- 1 . 4 5 8}$ | $\mathbf{0 . 0 0 1}$ | MCV coverage of $60-79 \%$ | $\mathbf{- 1 . 6 3 9}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of $75-84 \%$ | $\mathbf{- 2 . 0 4 2}$ | $\mathbf{0 . 0 0 0}$ | MCV coverage of $80-89 \%$ | $\mathbf{- 2 . 2 9 8}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of $85-94 \%$ | $\mathbf{- 2 . 6 2 3}$ | $\mathbf{0 . 0 0 0}$ | MCV coverage of $90-94 \%$ | $\mathbf{- 2 . 5 7 6}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of $>=95 \%$ | $\mathbf{- 2 . 9 7 7}$ | $\mathbf{0 . 0 0 0}$ | MCV coverage of $>=95 \%$ | $\mathbf{- 2 . 9 2 4}$ | $\mathbf{0 . 0 0 0}$ |
| ICD-8 coding system | 0.467 | 0.169 | ICD-8 coding system | 0.499 | 0.141 |
| ICD-9 coding system | 0.688 | 0.185 | ICD-9 coding system | 0.731 | 0.159 |
| ICD-10 coding system | 1.241 | 0.062 | ICD-10 coding system | $\mathbf{1 . 3 2 1}$ | $\mathbf{0 . 0 4 8}$ |
| Year | $\mathbf{- 0 . 1 1 9}$ | $\mathbf{0 . 0 0 0}$ | Year | $\mathbf{- 0 . 1 1 7}$ | $\mathbf{0 . 0 0 0}$ |
| Two doses of MCV | -0.374 | 0.168 | Two doses of MCV | -0.396 | 0.146 |
| Crude birth rate | $\mathbf{3 . 3 6 4}$ | $\mathbf{0 . 0 0 0}$ | Crude birth rate | $\mathbf{3 . 4 5 0}$ | $\mathbf{0 . 0 0 0}$ |
| Urban | $\mathbf{0 . 0 7 7}$ | $\mathbf{0 . 0 1 1}$ | Urban | $\mathbf{0 . 0 7 8}$ | $\mathbf{0 . 0 1 0}$ |
| Population density | -1.199 | 0.144 | Population density | -1.302 | 0.114 |
| Per-capita GDP | -0.380 | 0.497 | Per-capita GDP | -0.464 | 0.411 |
| Background mortality rate | 0.289 | 0.081 | Background mortality rate | 0.300 | 0.070 |
| Constant | 225.874 | 0.000 | Constant | 223.175 | 0.000 |

*Results shown for linear regression of logged under-5 mortality

For ease of interpretation, we computed the \% reduction compared to country-years with no MCV coverage implied by the coefficients at each MCV coverage levels as $100 *[1-\exp (\beta)]$, where $\beta$ is the regression coefficient for a particular coverage indicator. For example, with the best specification, MCV coverage of $85-94 \%$ produces a $93 \%$ reduction (i.e., $100 *[1-\exp (-2.623)]$ ) compared to a $92 \%$ reduction for a coverage level of $90-94 \%$ with the base case specification.

While the dependent variable is measles mortality in children under 5, the base case uses MCV coverage in 12-24 month-olds lagged by 1 year. It is therefore possible that vaccination coverage in prior years (i.e., vaccinated 2 year-olds that are now 4 year-olds) have an effect on under- 5 mortality as well. At the same time, if case fatality from measles is higher in younger children, including coverage levels for somewhat older children may attenuate the estimated relationship. We use lagged 5 -year average MCV coverage in an alternate model specification. In fact, we considered lags in two ways. First, we constructed averages based on all observations within a 5 -year range. For example, if data were only available for periods 2 years and 4 years prior, then only these two observations were used to construct the lagged average. Second, we constructed averages requiring that all MCV coverage values be present in the previous 5-year range. The first approach preserves sample size but makes the exact definition of the lagged average harder to interpret. The second approach maintains a clear definition of the lagged average, but loses sample size, and selects certain countries with longer and more continuous data series (see Table S1D above). Compared to the base-case specification, the main findings of the regression were similar in the alternative specifications, although the magnitude of the impact of MCV coverage above $80 \%$ attenuated under the alternatives (Table S3D).

| Table S3D. Comparison of regression coefficients: 5-year average MCV coverage vs. 1-year |
| :--- | :---: | :---: | :--- | :--- | :--- |
| lags |

\(\left.$$
\begin{array}{lcc|}\hline \begin{array}{c}\text { 5-Year Average MCV Coverage Specification } \\
\text { (any year in 5-year range without MCV coverage } \\
\text { causes observation to drop) } \mathbf{n}=\mathbf{7 9 1} \\
\text { Coefficient }\end{array} & \begin{array}{c}\text { P- } \\
\text { Independent variables }\end{array}
$$ \& <br>

\hline MCV coverage of 1-59\%\end{array}\right]\)| MCV coverage of 60-79\% | $\mathbf{- 1 . 1 9 0}$ | $\mathbf{0 . 0 0 4}$ |
| :--- | :---: | :---: |
| MCV coverage of 80-89\% | $\mathbf{- 1 . 7 1 1}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of 90-94\% | $\mathbf{- 2 . 5 3 5}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of >=95\% | $\mathbf{- 2 . 1 2 6}$ | $\mathbf{0 . 0 0 0}$ |
| ICD-8 coding system | $\mathbf{- 2 . 2 9 4}$ | $\mathbf{0 . 0 0 0}$ |
| ICD-9 coding system | 0.451 | 0.212 |
| ICD-10 coding system | -0.046 | 0.942 |
| Year | 0.781 | 0.314 |
| Two doses of MCV | -0.067 | 0.073 |
| Crude birth rate | -0.417 | 0.177 |
| Urban | $\mathbf{2 . 2 8 1}$ | $\mathbf{0 . 0 2 0}$ |
| Population density | $\mathbf{0 . 0 8 3}$ | $\mathbf{0 . 0 1 3}$ |
| Per-capita GDP | -1.423 | 0.135 |
| Background mortality rate | -0.741 | 0.288 |
| Constant | $\mathbf{1 . 5 3 7}$ | $\mathbf{0 . 0 0 3}$ |

*Results shown for linear regression of logged under-5 mortality

Computing percentage reductions in mortality under different coverage levels as above, we found that the two alternative 5-year lagged specifications predicted reductions in measles deaths at MCV coverage of $85-94 \%$ of $87 \%$ or $88 \%$ compared to $92 \%$ in the base case specification.

## Supporting Information S4. Sensitivity analysis: alternative regressions among country subsets defined by income or region

We estimated our main model specification on subsets of our data, confining our analysis to countries with higher GDPs or countries outside of less developed regions. For all countries in the analysis, we compared their year 2000 GDP per-capita (Penn World Tables 6.2, purchasing power parity, Laspeyres) to two thresholds $\$ \$ 7,000$ and $I \$ 5,000$. In a second set of regressions, we also excluded countries from Latin America or from Eastern Europe and the Former Soviet Union. We found that in wealthier countries in our sample (>I\$7,000), the impact of MCV coverage (especially above $80 \%$ ) attenuated slightly compared to the base case analysis (Table $\mathbf{S 4 A}$ ). By contrasting the MCV coefficients from the $>\mathrm{I} \$ 7,000$ and $>I \$ 5,000$ regressions with the base case, we see that the impact of vaccination was strongest in lower income countries, especially in the range between I $\$ 5,000$ and $\mathrm{I} \$ 7,000$.

Table S4A. Regression coefficients excluding countries with low per-capita GDP*

|  | Per-capita GDP <br> >I\$7,000 |  | Per-capita GDP <br> >IS5,000 |  | Base case |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :---: |
| Independent variables | Coefficient | P-value | Coefficient | P-value | Coefficient | P-value |
| MCV coverage of $1-59 \%$ | -0.317 | 0.404 | -0.274 | 0.466 | -0.236 | 0.523 |
| MCV coverage of $60-79 \%$ | $\mathbf{- 1 . 6 4 2}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 1 . 7 1 3}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 1 . 6 3 9}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of $80-89 \%$ | $\mathbf{- 2 . 0 5 7}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 3 7 2}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 2 9 8}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of $90-94 \%$ | $\mathbf{- 2 . 4 4 9}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 6 1 6}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 5 7 6}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of $>=95 \%$ | $\mathbf{- 2 . 6 8 3}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 9 4 8}$ | $\mathbf{0 . 0 0 0}$ | $-\mathbf{- 2 . 9 2 4}$ | $\mathbf{0 . 0 0 0}$ |
| ICD-8 coding system | 0.502 | 0.152 | 0.452 | 0.186 | 0.499 | 0.141 |
| ICD-9 coding system | 0.629 | 0.244 | 0.675 | 0.200 | 0.731 | 0.159 |
| ICD-10 coding system | $\mathbf{1 . 7 0 5}$ | $\mathbf{0 . 0 1 6}$ | 1.306 | 0.056 | $\mathbf{1 . 3 2 1}$ | $\mathbf{0 . 0 4 8}$ |
| Year | $\mathbf{- 0 . 1 2 0}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 0 . 1 1 8}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 0 . 1 1 7}$ | $\mathbf{0 . 0 0 0}$ |
| Two doses of MCV | $\mathbf{- 0 . 6 6 3}$ | $\mathbf{0 . 0 2 1}$ | -0.419 | 0.132 | -0.396 | 0.146 |
| Crude birth rate | $\mathbf{4 . 3 5 6}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{3 . 3 7 2}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{3 . 4 5 0}$ | $\mathbf{0 . 0 0 0}$ |
| Urban | $\mathbf{0 . 0 6 6}$ | $\mathbf{0 . 0 4 3}$ | $\mathbf{0 . 0 9 4}$ | $\mathbf{0 . 0 0 3}$ | $\mathbf{0 . 0 7 8}$ | $\mathbf{0 . 0 1 0}$ |
| Population density | -1.004 | 0.295 | $\mathbf{- 1 . 9 3 5}$ | $\mathbf{0 . 0 3 1}$ | -1.302 | 0.114 |
| Per-capita GDP | -0.160 | 0.827 | -0.452 | 0.444 | -0.464 | 0.411 |
| Background mortality rate | 0.136 | 0.440 | 0.297 | 0.074 | 0.300 | 0.070 |
| Constant | 223.247 | 0.000 | 225.573 | 0.000 | 223.175 | 0.000 |

*Results shown for linear regression of logged under-5 mortality

For ease of interpretation, we computed the $\%$ reduction compared to country-years with no MCV coverage implied by the coefficients at each MCV coverage levels as $100 *[1-\exp (\beta)]$, where $\beta$ is the regression coefficient for a particular coverage indicator. For example, the model estimated on countries with per-capita GDPs of $>\$ 7,000$ and $>\$ 5,000$ predict reductions in measles deaths of at MCV coverage of $85-94 \%$ of $91 \%$ and $93 \%$ respectively, compared to $92 \%$ with the base case specification.

It also appears that in Latin American and Eastern European and Former Soviet countries the impact of MCV was strongest (Table S4B), although differences were substantively negligible. Computing percentage reductions in mortality under different coverage levels as above, we found that the model estimated on countries excluding those in Latin America or excluding Eastern European Former Soviet countries each predicted reductions in measles deaths of $92 \%$ at MCV coverage of $90-94 \%$, equivalent to the predicted reduction in the base case specification.

Table S4B. Regression coefficients excluding countries from less developed regions

|  | Non-Latin America |  | Non-Eastern Europe <br> Former |  | Baviet case |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :---: |
| Independent variables | Coefficient | P-value | Coefficient | P-value | Coefficient | P-value |
| MCV coverage of 1-59\% | -0.793 | 0.046 | -0.126 | 0.737 | -0.236 | 0.523 |
| MCV coverage of $60-79 \%$ | $\mathbf{- 2 . 0 9 8}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 1 . 5 0 2}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 1 . 6 3 9}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of $80-89 \%$ | $\mathbf{- 2 . 4 5 8}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 1 5 3}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 2 9 8}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of 90-94\% | $\mathbf{- 2 . 5 0 3}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 5 8 2}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 5 7 6}$ | $\mathbf{0 . 0 0 0}$ |
| MCV coverage of >=95\% | $\mathbf{- 2 . 8 5 1}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 8 6 7}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{- 2 . 9 2 4}$ | $\mathbf{0 . 0 0 0}$ |
| ICD-8 coding system | 0.442 | 0.194 | 0.613 | 0.077 | 0.499 | 0.141 |
| ICD-9 coding system | 0.504 | 0.382 | 0.687 | 0.194 | 0.731 | 0.159 |
| ICD-10 coding system | $\mathbf{1 . 3 8 7}$ | $\mathbf{0 . 0 4 7}$ | $\mathbf{1 . 4 5 3}$ | $\mathbf{0 . 0 3 7}$ | $\mathbf{1 . 3 2 1}$ | $\mathbf{0 . 0 4 8}$ |
| Year | $\mathbf{- 0 . 1 1 7}$ | $\mathbf{0 . 0 0 2}$ | $\mathbf{- 0 . 0 8 9}$ | $\mathbf{0 . 0 0 4}$ | $\mathbf{- 0 . 1 1 7}$ | $\mathbf{0 . 0 0 0}$ |
| Two doses of MCV | 0.077 | 0.793 | $\mathbf{- 0 . 6 6 2}$ | $\mathbf{0 . 0 2 0}$ | -0.396 | 0.146 |
| Crude birth rate | $\mathbf{2 . 8 2 4}$ | $\mathbf{0 . 0 0 2}$ | $\mathbf{4 . 4 6 2}$ | $\mathbf{0 . 0 0 0}$ | $\mathbf{3 . 4 5 0}$ | $\mathbf{0 . 0 0 0}$ |
| Urban | $\mathbf{0 . 0 8 5}$ | $\mathbf{0 . 0 0 7}$ | 0.051 | 0.100 | $\mathbf{0 . 0 7 8}$ | $\mathbf{0 . 0 1 0}$ |
| Population density | -1.915 | 0.315 | -0.893 | 0.296 | -1.302 | 0.114 |
| Per-capita GDP | -0.279 | 0.669 | -1.017 | 0.140 | -0.464 | 0.411 |
| Background mortality rate | 0.592 | 0.296 | 0.249 | 0.132 | 0.300 | 0.070 |
| Constant | 223.647 | 0.003 | 168.935 | 0.003 | 223.175 | 0.000 |

In general, excluding poorer countries or those from less developed regions had relatively modest effects on our results. This analysis suggests that our base-case analysis may actually underestimate slightly the potential benefit of increasing MCV coverage in other parts of the world.

## Supporting Information S5. Sensitivity analysis: year linear trends versus year fixed effects

The main specification of the model includes calendar year as a linear trend. We find a significant linear trend in year with a slope of -0.117 (Reduction of $11.7 \%$ in death rate each year after 1960). As an alternative, we instead used year fixed effects (one dummy variable for each year). In the graph below, the coefficients for each year's fixed effect (and confidence intervals) are plotted. The coefficients can be thought of as the logarithm of the odds ratio and also connote a percent reduction from baseline since the model specification is log-linear. The slope of a line fit through the coefficients $(-0.1188)$ is highly concordant with the year linear trend that was defined prospectively and used in this main base case analysis - suggesting that long-term time patterns are generally captured with the linear trend. Other regression coefficients do not change substantially with the use of the year fixed effect specification (not shown).

Figure S5A. Comparison of year trend versus year fixed effects


## Supporting Information S6. Country-years with no observed measles deaths

For analyses of reductions in measles death rates, the rates were log-transformed (natural log). To deal with years with zero observed deaths (for which the log transformation would be undefined), we replaced the 0 with 0.1 times the minimum observed measles-specific death rate in children under 5 . As this was a prospectively-defined but arbitrary choice, we explored the effect of alternative replacement values on model results. We replaced country-years having 0 observed deaths with either the minimum observed measles death rate for children under 5 (i.e., a rate 10 times greater than in the main analysis) or else 0.01 times the minimum observed rate (i.e., a rate 10 times smaller than in the main analysis). We then estimated the model using these alternative outcome variables and compared the resulting coefficients for MCV coverage to the coefficients and $95 \%$ confidence intervals from the main analysis. Figure S6A shows the results of this sensitivity analysis. The alternative replacements do change the estimated effect, though the magnitude of the change falls within the $95 \%$ confidence intervals of the main analysis.

Another alternative would be to use a statistical model for count data such as a Poisson panel regression or a negative binomial panel regression with either fixed or random effects. As Poisson panel regressions are a special case of negative binomial panel regressions in which the mean and variance are assumed to be equal, we estimated negative binomial panel regressions with both fixed and random effects (Table S6B). Whereas in the main analysis, MCV coverage of $80 \%$ or greater had substantially greater impact than lower levels of coverage, in these alternative models, the effect was more continuous across the coverage levels. Furthermore, at very high coverage levels the reduction compared to country-years with no MCV coverage was estimated to be approximately $80 \%$. The magnitude of the estimated benefit was most consistent with replacing years with no observed deaths with the minimum observed death rates ( 10 times larger than the main analysis) (Figure S6A).

Figure S6A. Comparison of regression models


Table S6B. Alternative models: negative binomial panel regressions*

|  | Conditional country fixed effects |  |  | Country random effects |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Incidence- <br> rate <br> Ratio | 95\% CI |  | Incidence- <br> rate <br> Ratio |  |  |
| MCV coverage of 1-59\% | 0.42 | 0.35 | 0.52 | 0.41 | 0.34 | 0.50 |
| MCV coverage of 60-79\% | 0.28 | 0.21 | 0.38 | 0.27 | 0.20 | 0.36 |
| MCV coverage of 80-89\% | 0.22 | 0.15 | 0.33 | 0.22 | 0.15 | 0.32 |
| MCV coverage of 90-94\% | 0.25 | 0.17 | 0.37 | 0.24 | 0.16 | 0.36 |
| MCV coverage of $>=95 \%$ | 0.20 | 0.13 | 0.30 | 0.19 | 0.12 | 0.29 |
| ICD-8 coding system | 1.26 | 1.05 | 1.50 | 1.28 | 1.07 | 1.53 |
| ICD-9 coding system | 1.27 | 0.86 | 1.85 | 1.30 | 0.89 | 1.92 |
| ICD-10 coding system | 1.15 | 0.61 | 2.17 | 1.15 | 0.61 | 2.17 |
| Year | 0.93 | 0.91 | 0.94 | 0.93 | 0.92 | 0.95 |
| Two doses of MCV | 0.83 | 0.60 | 1.14 | 0.80 | 0.59 | 1.10 |
| Crude birth rate | 1.28 | 0.94 | 1.73 | 1.29 | 0.97 | 1.74 |
| Urban | 0.98 | 0.97 | 0.99 | 0.98 | 0.97 | 0.99 |
| Population density | 1.79 | 1.55 | 2.07 | 1.66 | 1.45 | 1.90 |
| Per-capita GDP | 0.79 | 0.63 | 0.98 | 0.73 | 0.59 | 0.91 |
| Background mortality rate | 0.98 | 0.90 | 1.07 | 0.98 | 0.91 | 1.06 |

* 958 country-years for 42 countries (fixed effects model); 980 country-years for 44 countries (random effects model). Incidence compared to country-years with no measles coverage.

