

EM Algorithm

In what follows, we give the procedure for estimating parameters in Θ within the EM algorithm framework. In the E-step, the posterior expectation of z_{ij} is evaluated as

$$P_{ij} = E[z_{ij}|\Theta, \mathbf{y}_i] = Pr[z_{ij} = 1|\Theta, \mathbf{y}_i] = \frac{\omega_j f_{ij}(\mathbf{y}_i; \boldsymbol{\mu}_{ij}, \Sigma_i)}{\sum_{j'=1}^J \omega_{j'} f_{ij'}(\mathbf{y}_i; \boldsymbol{\mu}_{ij'}, \Sigma_i)}. \quad (1)$$

In the M-step, closed form solutions exist for $\boldsymbol{\omega}$ and the parameters in Θ_{μ_j} except for τ and σ^2 . Since $\omega_J = 1 - \sum_{j=1}^{J-1} \omega_j$, it is easy to show that for $j = 1, \dots, J-1$,

$$\frac{\partial \log L_c(\Theta|\mathbf{y})}{\partial \omega_j} = \sum_{i=1}^n \frac{P_{ij}}{\omega_j} - \sum_{i=1}^n \frac{P_{iJ}}{1 - \sum_{j'=1}^{J-1} \omega_{j'}}.$$

By setting the above equation to zero and solving it for ω_j , we have

$$\omega_j = \frac{\sum_{i=1}^n P_{ij}(1 - \sum_{j'=1}^{J-1} \omega_{j'})}{\sum_{i=1}^n P_{iJ}}. \quad (2)$$

Calculating the sum of both sides of Equation (2) over $j = 1, \dots, J-1$, we get

$$1 - \omega_J = \frac{\sum_{i=1}^n \sum_{j=1}^{J-1} P_{ij} \omega_J}{\sum_{i=1}^n P_{iJ}}. \quad (3)$$

Since $\sum_{j=1}^{J-1} P_{ij} = 1 - P_{iJ}$, solving (3) for ω_J , it can be shown that

$$\hat{\omega}_J = \sum_{i=1}^n P_{iJ}/n \quad (4)$$

Plugging (4) back into (2), we have $\hat{\omega}_j = \sum_{i=1}^n P_{ij}/n$.

Suppose the gene expression trajectory is approximated by the first K orders of the Fourier series, then $\Theta_{\mu_j} = (\mathbf{c}_j, \tau_j)$, where $\mathbf{c}_j = (\alpha_{0j}, \alpha_{1j}, \beta_{1j}, \dots, \alpha_{Kj}, \beta_{Kj})$. We have

$$\frac{\partial \log L_c(\Theta|\mathbf{y})}{\partial \mathbf{c}_j} = \left[\frac{\partial \log L_c(\Theta|\mathbf{y})}{\partial \boldsymbol{\mu}_{ij}} \right] \left[\frac{\partial \boldsymbol{\mu}_{ij}}{\partial \mathbf{c}_j} \right]. \quad (5)$$

The parameter $\mathbf{c}_j$ can be updated by setting (5) to zero. Since

$$\frac{\partial \log L_c(\Theta|\mathbf{y})}{\partial \boldsymbol{\mu}_{ij}} = \sum_{i=1}^n P_{ij} (\mathbf{y}_i - \boldsymbol{\mu}_{ij})^T \Sigma_i^{-1}$$

and $\frac{\partial \boldsymbol{\mu}_{ij}}{\partial \mathbf{c}_j} = D_i(\tau_j)$, where

$$D_i(\tau_j) = \begin{pmatrix} 1 & \cos(\frac{2\pi t_{i1}}{\tau_j}) & \sin(\frac{2\pi t_{i1}}{\tau_j}) & \dots & \cos(\frac{2\pi K t_{i1}}{\tau_j}) & \sin(\frac{2\pi K t_{i1}}{\tau_j}) \\ 1 & \cos(\frac{2\pi t_{i2}}{\tau_j}) & \sin(\frac{2\pi t_{i2}}{\tau_j}) & \dots & \cos(\frac{2\pi K t_{i2}}{\tau_j}) & \sin(\frac{2\pi K t_{i2}}{\tau_j}) \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & \cos(\frac{2\pi t_{im_i}}{\tau_j}) & \sin(\frac{2\pi t_{im_i}}{\tau_j}) & \dots & \cos(\frac{2\pi K t_{im_i}}{\tau_j}) & \sin(\frac{2\pi K t_{im_i}}{\tau_j}) \end{pmatrix},$$

we have

$$\hat{\mathbf{c}}_j = \left[\sum_{i=1}^n P_{ij} D_i(\tau_j)^T \Sigma_i^{-1} D_i(\tau_j) \right]^{-1} \left[\sum_{i=1}^n P_{ij} \mathbf{y}_i^T \Sigma_i^{-1} D_i(\tau_j) \right].$$

Since the analytical form of the inverse of Σ_i is not available, we use the recursive method proposed by Haddad (2004) to calculate the inverse matrix of ARMA(p, q) through its association with ARMA(p, q-1).

We can write $\Sigma_i = \sigma^2 R_i$, where R_i is the correlation matrix that is entirely determined by the ARMA parameters $\varphi_1, \dots, \varphi_p, \theta_1, \dots, \theta_q$. The variance σ^2 can be updated by

$$\hat{\sigma}^2 = \frac{\sum_{i=1}^n \sum_{j=1}^J P_{ij} (\mathbf{y}_i - \mu_{ij})^T R_i^{-1} (\mathbf{y}_i - \mu_{ij})}{\sum_{i=1}^n m_i}. \quad (6)$$

Again R_i^{-1} can be calculated by the method of Haddad (2004).

Because there are no closed form solutions for τ_j and ARMA parameters $\varphi_1, \dots, \varphi_p$ and $\theta_1, \dots, \theta_q$, their estimates are updated using one-step Newton-Raphson method within each iteration. In particular, in the $(\nu + 1)$ -th iteration, τ_j can be updated by

$$\tau_j^{\nu+1} = \tau_j^\nu - \frac{\frac{\partial}{\partial \tau_j} \log L_c(\Theta | \mathbf{y})|_{\Theta=\Theta^\nu}}{\frac{\partial^2}{\partial \tau_j^2} \log L_c(\Theta | \mathbf{y})|_{\Theta=\Theta^\nu}}, \quad (7)$$

where

$$\frac{\partial}{\partial \tau_j} \log L_c(\Theta | \mathbf{y}) = \sum_{i=1}^n P_{ij} (\mathbf{y}_i - \mu_{ij})^T \Sigma_i^{-1} \delta_{ij}$$

with δ_{ij} being a $m_i \times 1$ vector whose components

$$\delta_{ijl} = \sum_{k=1}^K \left[\alpha_{kj} \sin \left(\frac{2\pi k t_l}{\tau_j} \right) \frac{2\pi k t_l}{\tau_j^2} - \beta_{kj} \cos \left(\frac{2\pi k t_l}{\tau_j} \right) \frac{2\pi k t_l}{\tau_j^2} \right],$$

and

$$\frac{\partial^2}{\partial \tau_j^2} \log L_c(\Theta | \mathbf{y}) = \sum_{i=1}^n \left[-P_{ij} \delta_{ij}^T \Sigma_i^{-1} \delta_{ij} + P_{ij} (\mathbf{y}_i - \mu_{ij})^T \Sigma_i^{-1} \frac{\partial^2}{\partial \tau_j^2} \mu_{ij} \right].$$

Similarly, the parameters $\varphi_1, \dots, \varphi_p$ and $\theta_1, \dots, \theta_q$ can be updated by the one-step Newton-Raphson method outlined above. However, there are no analytical forms of the first and the second derivatives of the expected complete data log-likelihood with respect to the φ 's and θ 's, we use the numerical differentiation method to calculate these quantities (Zeng and Cai, 2005). To ease the presentation of the method, denote the $(p + q)$ dimensional vector $\psi = (\varphi_1, \dots, \varphi_p, \theta_1, \dots, \theta_q)$. The first and the second derivatives with respect to the κ -th component in ψ are approximated, respectively, by

$$\frac{E[\log L_c(\Theta_{-\psi}, \psi + h_n e_\kappa | \mathbf{y})] - E[\log L_c(\Theta | \mathbf{y})]}{h_n}, \quad (8)$$

and

$$\frac{E[\log L_c(\Theta_{-\psi}, \psi + h_n e_\kappa | \mathbf{y})] - 2E[\log L_c(\Theta | \mathbf{y})] + E[\log L_c(\Theta_{-\psi}, \psi - h_n e_\kappa | \mathbf{y})]}{h_n^2}, \quad (9)$$

where we use E to represent the posterior expectation of the complete data log-likelihood with respect to w_{ij} , $\Theta_{-\psi}$ denotes the parameters in Θ other than ψ , the $(p+q)$ vector e has unity length with the κ -th component set to 1, and h_n is the bandwidth chosen by the investigator. When h_n is small enough, the numerical differentiation approximates the true derivatives adequately. On the other hand, if h_n is too small, the random errors from the numerical computation may deteriorate the results.